# Mathematical Statistics Preliminary Examination 

Statistics Group, Department of Mathematics and Statistics, Auburn University
August 21, $2012 \quad 9: 00 \mathrm{AM}-12: 00 \mathrm{PM}$
Solve any five problems out of the seven problems. The maximum possible score is 50 points. Each question is worth 10 points. If you work out more than five problems, your score is the sum of five highest points.

## Name:

## Questions:

1. State and prove the Cramér-Rao inequality. What are two uses of this theorem?
2. Let $X_{1}, \ldots, X_{n}$ be a random sample from the density $f_{\theta}(x)=\theta f_{1}(x)+(1-\theta) f_{2}(x)$, where $f_{1}$ and $f_{2}$ are two different densities and $\theta \in(0,1)$ is unknown.
(a) Provide a necessary and sufficient condition for the likelihood equation to have a unique solution and show that if there is a solution, then it is the MLE of $\theta$.
(b) Derive the MLE of $\theta$ when the likelihood equation has no solution.
3. Suppose that $T_{n}$ has a $\chi^{2}$ distribution of $n$ degrees of freedom. Show that $\sqrt{2 T_{n}}-\sqrt{2 n}$ converges in distribution to $N(0,1)$. Hint: A chi-squared random variable is the sum of squared standard Normal random variables.
4. Suppose that $X_{1}, \ldots, X_{n}$ are independent and

$$
P\left(X_{i}=k\right)=\frac{P\left(\operatorname{Binomial}\left(m_{i}, \theta\right)=k\right)}{P\left(\operatorname{Binomial}\left(m_{i}, \theta\right)>0\right)}=\frac{\binom{m_{i}}{k} \theta^{k}(1-\theta)^{m_{i}-k}}{1-(1-\theta)^{m_{i}}}
$$

for $k=1, \ldots, m_{i}$. The numbers $m_{1}, \ldots, m_{n}$ are known integers with $m_{i} \geq 1$. That is, $X_{i}$ has the conditional distribution of a $\operatorname{Binomial}\left(m_{i}, \theta\right)$ random variable given that it is not 0 .
(a) Find a complete and sufficient statistic.
(b) Assume that all the $m_{i}$ are the same value $m$. Use the random variable $Y=$ $I\left(X_{i}=m\right)$ to give a formula for the UMVUE of $\theta^{m} /\left[1-(1-\theta)^{m}\right]$. You are not expected to simplify the formula.
(c) For the case $n=2$ and $m_{1}=m_{2}=m=3$, what is the value of the estimate in the previous part if $X_{1}=3$ and $X_{2}=1$ ?
5. Suppose $Y$ is a random variable with $\operatorname{pdf} g(y), U$ is a random variable with a uniform $(0,1)$ distribution, and $Y$ and $U$ are independent. Let $f(x)$ be a pdf. Suppose that there is a constant $M$ such that $f(x) \leq M g(x)$ for all $x \in \mathbb{R}$. Define a random variable $X$ as $X=Y$ if $U \leq \frac{f(Y)}{M g(Y)}$. Show that $X$ has pdf $f(x)$.
6. Suppose the random variables $C, T$, and $Z$ are related as follows:

$$
P(T \geq t \mid Z)=e^{-Z \theta t}, \quad P(C \geq c \mid Z)=e^{-Z \rho c}
$$

and $T$ and $C$ are independent given $Z$. Assume further that $Z$ has a gamma distribution with mean 1 and variance $\phi^{-1}$; that is, it has the moment generating function $M_{Z}(t)=$ $(1-t / \phi)^{-\phi}$. Show that the joint distribution function of $(T, C)$ can be found as

$$
P(T \geq t, C \geq c)=\left(1+\frac{\theta t}{\phi}+\frac{\rho c}{\phi}\right)^{-\phi} .
$$

7. Find the likelihood ratio test (LRT) of

$$
H_{0}: \alpha \leq 0 \quad \text { vs } \quad H_{1}: \alpha>0
$$

based on a sample $X_{1}, \ldots, X_{n}$ from the two-parameter exponential distribution with pdf

$$
f(x ; \alpha, \beta)=\frac{1}{\beta} \exp \left(-\frac{x-\alpha}{\beta}\right) I_{[\alpha, \infty)}(x)
$$

where both $-\infty<\alpha<\infty$ and $\beta>0$ are unknown.

